

# Modelling spousal mortality dependence: evidence of heterogeneities and implications

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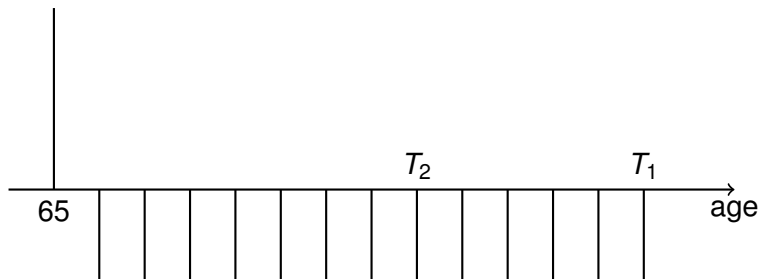


# INTRODUCTION

# Motivation

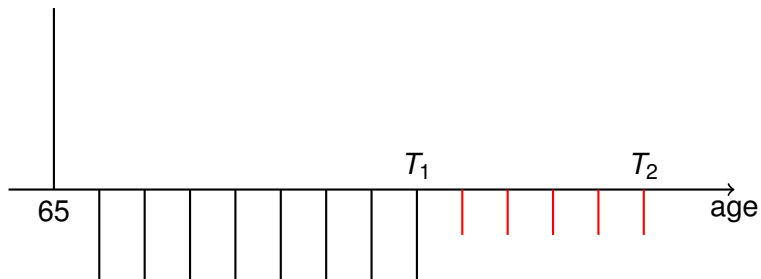
- ▶ It has been documented that the two spouses have positive mortality dependence.
- ▶ Many pension funds/annuity providers offer reversionary benefit.
- ▶ The aim of this paper is to model such a dependence, and discuss its impact on pricing.

## Cash flow of a joint annuity: Case A



**Figure:** In Case A, the principal beneficiary dies after the spouse ( $T_1 > T_2$ ). No additional benefit for the spouse.

## Case B



**Figure:** In Case B, the principal beneficiary dies before the spouse ( $T_1 < T_2$ ). Spouse continues to receive a reduced annuity.

## Different types of dependencies

The mortality jump can be due to two reasons:

- ▶ The broken-heart syndrome, that is a causal effect.
- ▶ Spurious risk correlation due to similar risk factors, i.e. observed and unobserved heterogeneities.

Current literature usually neglects the heterogeneities.

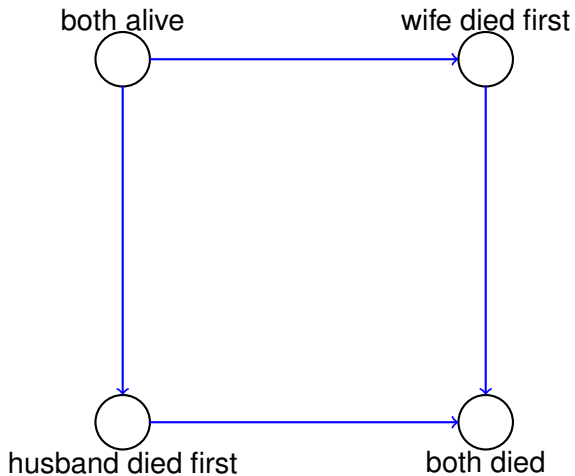
It is important to disentangle these different types of dependencies [see Gourieroux and Lu (2015 IME)].

# Outline

- ▶ We estimate a model which distinguishes them.
- ▶ We find that heterogeneities explain a significant proportion of the dependency. These heterogeneities can be either observed or unobserved.
- ▶ We find important over/under estimations of the standard model depending on the characteristics of the couple.

## THE MODEL

## The status of the couple



## Before the first death: spurious risk correlation

The mortality intensities are:

$$\theta_1(t|U, V, T_1 > t, T_2 > t, X, c_1, c_2) = \lambda_1(t + c_1) \exp(\beta_1' X) U,$$

$$\theta_2(t|U, V, T_1 > t, T_2 > t, X, c_1, c_2) = \lambda_2(t + c_2) \exp(\beta_2' X) V,$$

where:

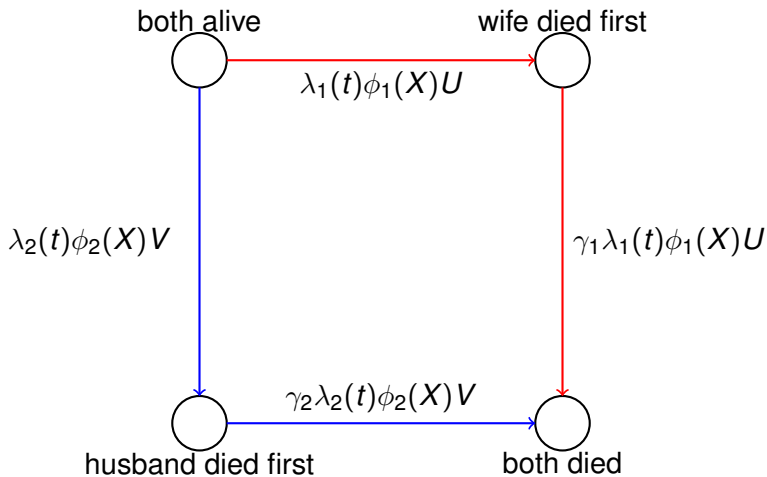
- ▶  $T_1, T_2$  are duration,  $c_1, c_2$  are ages at inception;
- ▶  $X$  is the covariate(s), including ages at inception and dates of birth;
- ▶ Thus there can be duration effects, age effects, and even cohort effects;
- ▶  $U, V$  are the unobserved individual heterogeneities; they introduce *spurious dependence* between  $T_1, T_2$ ;
- ▶  $\lambda_1, \lambda_2$  are baseline intensities;

This is called mixed proportional hazard competing risks model.

## After the first death: the broken-heart syndrome

- ▶ After the death of male (resp. female), the mortality of the spouse is multiplied by a constant  $\gamma_1$  (resp.  $\gamma_2$ ).
- ▶ These parameters capture the real duration dependence, the "contagion".
- ▶ Usually  $\gamma_1 \neq \gamma_2$ .

## (Conditional) transition intensities



The parameters of interest are:

- ▶ Baseline intensities  $\lambda_1, \lambda_2$ , which we specify as exponential functions.
- ▶ The distribution of  $(U, V)$ , which we specify in a semi-parametric way.
- ▶ Regression coefficients  $\beta_1, \beta_2$ .
- ▶ Jump parameters  $\gamma_1, \gamma_2$ .

# The data

- ▶ A French joint annuity portfolio kindly provided by Scor.
- ▶ Population size: 27000 couples.
- ▶ Observation period: 2005-2013.
- ▶ Available information: ages at inception, amount of the annuity, dates of birth, potential dates of death (with truncations and censoring).
- ▶ The model is estimated by maximum likelihood. See the paper for details.

## EMPIRICAL RESULTS

## What we find

- ▶ Asymmetric jumps:  $\gamma_1 = 1.53 > \gamma_2 = 1.28$ .
- ▶ The unobserved heterogeneities of the two spouses are highly positively correlated.
- ▶ They account for a significant portion of the dependency that cannot be explained by the broken-heart syndrome.
- ▶ Moreover the parameters  $\gamma_1, \gamma_2$  are over-estimated if unobserved heterogeneities are omitted.

## Comparison of three models

Model	unobserved heterogeneity	mortality jump
1	Yes	Yes
2	No	Yes
3	No	No

Table: Summary of the three models

We compare their forecasts of the following quantities:

- ▶ Residual life expectancies.
- ▶ The price of Contract 1 (resp. 2), for a male (female) principal beneficiary, with a  $1/3$  discount for the spouse if she/he survives longer.

Case 1: Male is aged 60, female 65

Model	male residual life expectancy	female residual life expectancy	Price Contract No. 1	Price Contract No. 2
1	23.06	24.47	25.56	26.26
2	22.86	24.28	25.35	26.05
3	23.53	23.02	25.93	25.68

Case 2: Male is aged 70, female 65

Model	male residual life expectancy	female residual life expectancy	Price Contract No. 1	Price Contract No. 2
1	15.37	23.93	20.23	24.51
2	15.24	23.64	20.02	24.23
3	15.22	23.02	20.04	23.94

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# Discussions

- ▶ The standard model that omits dependencies lead to huge error.
- ▶ This error can be either positive, or negative, depending on the age of both spouses.
- ▶ Omitting unobserved heterogeneities leads to under-estimation of the life expectancies.

## CONCLUSION

- ▶ We estimated a model which captures the joint dependency of the two lifetimes.
- ▶ We found evidence of heterogeneities on the mortality rates, besides the broken-heart syndrome. We assessed their impacts.
- ▶ Finally, the findings are also useful for the pricing of single life insurance.

THANKS FOR YOUR ATTENTION.  
QUESTIONS/SUGGESTIONS?